

Problems of the gonad index and what can be done: analysis of the purple sea urchin *Strongylocentrotus purpuratus*

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Abstract The gonad index, GI, is widely used as a measure of changes in reproductive state. There are, however, problems with its use because it is based on the implicit assumption of an isometric relationship between gonad size and some measure of total size. If, for example, gonad weight and total weight are used, the exponent for an allometric relationship usually is ignored and hence assumed to be 1.0. It is further assumed that this exponent is fixed for all states of the reproductive cycle and that gonads begin to develop at size = 0. Data for the purple sea urchin *Strongylocentrotus purpuratus* at Gregory Point, Oregon, USA, gathered over a period of 31 months showed that these assumptions cannot be supported. The relationship is better modeled with a function that (1) takes into account size of initial gonad production and (2) allows allometric exponents that vary with site or season. Thus, a better approach is to use a wide range of sizes to estimate size when gonads begin to develop and then, with this

correction, ANCOVA to test for differences of gonad size among samples. Gonad changes at Gregory Point were estimated using fixed sizes of 5 cm diameter and 60 g total weight. Publishing means for X and Y , the standard error of the estimate, R^2 , and slope for each regression are shown to be sufficient to compare our results with results across studies.

Introduction

Describing morphological changes in body parts with ratios is pervasive in biology and has been used for decades. Criticism of ratios or body indices also is of long standing (e.g., Atchley et al. 1976; deVlaming et al. 1982) and consensus of critics is that ANCOVA is a superior approach to analysis (e.g., Packard and Boardman 1987, 1999; Raubenheimer and Simpson 1992). Comparison of analyses using ratios and ANCOVA has shown that errors due to the statistical analysis of ratio-based indices may be more serious than generally appreciated (Beaupre and Dunham 1995). Nevertheless, ratios are intuitive and easy to calculate whereas “...use and interpretation of ANCOVA requires a greater statistical sophistication ...” (Liermann et al. 2004). With the nearly universal availability of statistical software, however, ANCOVA is readily available and not difficult to use with the sorts of data sets associated with changes in sizes of body parts. Work we present here focuses on use of a gonad index in sea urchins but the problem is very general because dissections of many organisms are done for various reasons and indices continue to be the summary of choice.

The gonad index, gonosomatic index, or gonadosomatic index, GI, has taken a variety of forms for marine invertebrates (Table 1) including wet and dry weights (e.g.,

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Table 1 Examples of ways a gonad index, GI, has been calculated for different marine invertebrates and usually expressed as percent; this list is not exhaustive

Invertebrate group	References
<i>Echinodermata</i>	
Echinoidea	
Gonad wt wet/total wt wet	Lawrence et al. (1965), Giese (1966), Pearse (1968), Vernon et al. (1993), Vadas et al. (2000), McCarthy and Young (2002), Bigatti et al. (2004), Lester et al. (2007)
Gonad vol/total wt wet	Bennett and Giese (1955), Giese et al. (1959)
Gonad wt dry/total wt dry	Excluding body fluid and gut contents. Gonor (1972) Including gut contents. Laegdsgaard et al. (1991), Raymond and Scheibling (1987), Guillou and Lumingas (1999)
Gonad vol/test vol	Moore et al. (1963), McPherson (1965), Lewis and Storey (1984)
Gonad vol/test diameter	Lewis (1958)
Gonad vol/test diameter ³	Lewis (1966)
Gonad vol/body wt dry without gut contents	Lessios (1981)
4 gonads wt dry/body wt dry	Body is without gonads and gut, Lozano et al. (1995)
Gonad wt wet/drained body wt wet	Kramer and Nordin (1975), Bernard (1977)
Gonad wt wet/eviscerated test wt wet	Kelly et al. (2001), MacCord and Ventura (2004)
Gonad wt dry/test diameter ³	Guettaf et al. (2000)
Gonad wt dry/test diameter ² height	Ebert (1968)
Gonad wt dry/eviscerated test wt dry	Guillou and Lumingas (1998)
Ophiuroidea	
Gonad wt dry/eviscerated body wt dry	“Body” excludes arms, gut, and gonads; Bourgoin and Guillou (1990)
Gonad wt dry/total weight dry	Lefebvre et al. (1999)
Asteroidea	
Gonad wt wet/total wt wet	Lawrence (1973), Boivin et al. (1986), Chen and Chen (1992), Byrne et al. (1997)
Gonad wt dry/eviscerated body wt dry	“Eviscerated” = no gonads or pyloric caeca; Barker and Xu (1991), Byrne (1992)
Gonad wt dry/body wt dry (without gonads)	Nojima (1979)
Gonad vol/eviscerated wt wet	Without gonad and pyloric caeca; Town (1980), Nichols and Barker (1984)
Holothuroidea	
Gonad wt wet/body wall wt wet	Chao et al. (1995)
Gonad vol/body wall wt wet	Body weight is without viscera or body fluid; Cameron and Fankboner (1986)
Gonad wt wet/gutted body wt wet	Gutted body is whole animal minus gut and gonad; Gutt et al. (1992)
Gonad wt wet/total body drained of fluid	Ong Che and Gomez (1985)
<i>Arthropoda</i>	
Crustacea	
Gonad wt dry/carapace width	Diez and Lovrich (2010)
Gonad wt wet $\times 10^5/L^3$ (L = carapace length)	Pérez-González et al. (2009)
Gonad wt wet/carapace length	Quackenbush and Herrkind (1981)
Cnidaria	
Gonad wt dry/total wt dry	Chadwick-Furman et al. (2000)
Gonad vol/total wt wet	Ford (1964)
<i>Mollusca</i>	
Cephalopoda	
Nidamental gland or testis (mm)/mantle (mm)	Costa and Fernandes (1993)
Gonad wt wet/total body wt wet	Hatfield and Murray (1999)

Table 1 continued

Invertebrate group	References
Bivalvia	
Gonad wt wet/total tissue wt wet	Hines (1979)
Mantle wt dry/(body-mantle) wt dry	Nicastro et al. (2010)
Gastropoda	
Gonad wt wet/somatic tissue wt wet	Liu (1994)
Annelida	
Gonad energy/total body energy in joules	Olive et al. (1997)

Lawrence et al. 1965; Gonor 1972), volumes (e.g., Moore et al. 1963), linear measurements (e.g., Costa and Fernandes 1993) and units of energy (e.g., Olive et al. 1997). Mixtures of linear and weight measurements may or may not include exponents. For example, in studies of Crustacea, carapace width or length has been used without an exponent (Diez and Lovrich 2010) or with an assumed exponent of 3 (Pérez-González et al. 2009). With bivalves, the mantle may or may not be included in the denominator (Hines 1979; Nicastro et al. 2010).

For ophiuroids, body weight may mean total (Lefebvre et al. 1999) or excluding arms, gut, and gonads (Bourgoin and Guillou 1990). Similarly in asteroids, body weight may mean total (Boivin et al. 1986; Chen and Chen 1992; Byrne et al. 1997) or eviscerated with gonads removed (Nojima 1979) or both gonads and the pyloric caeca removed (Barker and Xu 1991; Byrne 1992). Measures could be either wet or dry. Gonad indices for holothurians generally use the body drained of fluid (e.g., Ong Che and Gomez 1985) but may also include removal of gut and gonad (e.g., Cameron and Fankboner 1986; Gutt et al. 1992).

In studies of sea urchins (echinoids), measurements have been of whole body weight and gonad volume (Bennett and Giese 1955), volume for both gonads and body (Moore et al. 1963), dry weights sometimes excluding gut contents (Gonor 1972) and sometimes including gut contents (Laegdsgaard et al. 1991). Sometimes only some of the five gonads have been used (Lozano et al. 1995; Byrne et al. 1998). When linear and weight measurements have been mixed, diameter has been cubed (e.g., Lewis 1966; Guettaf et al. 2000), sometimes not (Lewis 1958), and may include height as well as diameter (Ebert 1968). Finally, in some studies, body wet weight was used following evisceration or draining of body fluid (Kramer and Nordin 1975; Kato and Schroeter 1985; Guillou and Lumingas 1998; Kelly et al. 2001).

Over the years, analysis for sea urchins has settled on using wet weight of the gonad divided by the total wet weight multiplied by 100% (e.g., Hughes et al. 2006; Kirby et al. 2006; Lester et al. 2007; Lau et al. 2009), which is

$$GI = \text{gonad}/\text{total} \times 100\%. \quad (1)$$

The general form of the gonad index (GI), however, is an allometric equation (power function) of gonad size (G) and some measure of total body size (S)

$$G = \alpha S^\beta \quad (2)$$

and, with the assumption that $\beta = 1.0$, α is what Byrne et al. (1998) have called the “gonad retrieval rate”. The parameter α is the slope of a linear regression of gonad weight as a function of total weight with a 0-intercept. Rearrangement of Eq. 2 is

$$\alpha = G/S^\beta \quad (3)$$

and α is the gonad index, GI; the “gonad retrieval rate” is the same as Eq. 1. Equation 3 is the form of all of the examples in Table 1.

The first problem is that for Eq. 3 to be the same as Eq. 1, the exponent β must be equal to 1.0. If size, S , is a linear measurement the exponent β may be arbitrarily assigned the value 3 (e.g., Lewis 1966; Guettaf et al. 2000; Pérez-González et al. 2009) or may be used without an exponent, which assumes that β is 1.0 (Diez and Lovrich 2010). Damián et al. (2009) calculated β then fixed the value for an ANCOVA with a restricted set of size classes.

A second problem is that gonads begin to develop at a size that is not equal to zero and so Eq. 2 requires an additional parameter

$$G = \alpha(S - C)^\beta \quad (4)$$

where C is the size when gonads begin to develop. In general, this problem is not addressed (e.g., Diez and Lovrich 2010). Equation 4 has, however, been used to estimate C for several sea urchin species (Ebert and Russell 1994; Russell 1998; Ebert 1998, 2008, 2010b; Selden et al. 2009) but interest in the general problem of change in relative growth of gonads is of long standing (e.g., Robb 1929).

Additional problems with Eq. 1 center on the use of ratios when the independent variable appears as X but also is used in Y (cf. Atchley et al. 1976) as is the case where a body index is plotted as a function of total size to show a

size-dependent relationship (e.g., Gonor 1972). “Part-whole correlation” (Christians 1999) where total weight includes gonad weight, and the total weight changes as gonad weight changes, is not a problem in sea urchin studies: total weight is independent of gonad weight because of the inverse relationship between body fluid and gonad weight. This inverse relationship, however, is a problem with the calculation of a gonad index, GI, using drained weight (Kramer and Nordin 1975; Bernard 1977).

There are various uses of the GI with the most basic being a simple description of a periodic, usually annual, reproductive cycle (e.g., Bennett and Giese 1955; Lawrence et al. 1965; Pearse 1968; Lewis and Storey 1984; Bigatti et al. 2004). Although not certain, the GI probably is suitable if this is the only goal of the study. When comparisons of the GI using additional statistical analyses are made across sites (e.g., Guillou and Lumingas 1999; Lester et al. 2007), sizes (e.g., Gonor 1972), or seasons (e.g., Laegdsgaard et al. 1991) failure to use an appropriate allometric model may influence conclusions (cf. Beaupre and Dunham 1995).

Our goal is to show problems associated with application of a body index, and the gonad index in particular. We also demonstrate the use of ANCOVA in analysis of a 31-month study of gonad changes in the US west coast purple sea urchin *Strongylocentrotus purpuratus* and finally show how analysis in this manner permits use of the results in this paper to test for differences in future studies. In this regard, it is a problem of selecting the minimum number of summary statistics that need to be published in order to do a statistical analysis without access to raw data although raw data always will be best.

Methods

Purple sea urchins (*Strongylocentrotus purpuratus*) were collected monthly during low tides at Gregory Point, Oregon (43° 20' N; 124° 22' W) from January 2007 through July 2009. Test diameter was measured with digital calipers, total wet weight determined, and gonads removed and weighed wet. A range of sizes was collected each month and during the entire study spanned 2.5–9.2 cm for test diameter and 8.3–298 g for total wet weight.

The first step in analysis was to estimate the parameter C for both diameter and total wet weight in Eq. 4, which was done by nonlinear regression (NONLIN in SYSTAT 2004). Samples were selected that had individuals sufficiently small that estimates of C were less than 1 cm or 1 g total wet weight and included some from Gregory Point, Boiler Bay, Yaquina Head and Sunset Bay in Oregon and Arena Cove and Yankee Point in California. The model used was

$$\ln G = A + \beta \ln(S - C) \quad (5)$$

where G is gonad wet weight (g), A is $\ln \alpha$, and S is either diameter (cm) or total wet weight (g). A logarithmic transformation was used to minimize the influence of increasing variation in gonad sizes of large individuals on the estimate of C .

Following the determination of C , the next step was to test whether a single exponent, β , was appropriate for all monthly samples, which was done by an ANCOVA using the general linear model (GLM) in SYSTAT. There are 31 monthly samples and all could have the same value of β or β could change as gonads increase and decrease in weight. The model can be thought of as making A and β in Eq. 5 linear functions of the sample number (month) that goes from January 2007 (sample #1) to July 2009 (sample #31). Sample number was treated as a descriptive factor in the analysis and not ordinal. Samples could have been called a , b , c , etc. Including “sample” in the analysis is

$$A = b_1 \text{sample} + c_1 \quad (6)$$

$$\beta = b_2 \text{sample} + c_2 \quad (7)$$

Substituting Eqs. 6 and 7 into Eq. 5 is

$$\ln G = b_1 \text{sample} + c_1 + (b_2 \text{sample} + c_2) \times \ln(S - C) \quad (8)$$

or

$$\ln G = b_1 \text{sample} + c_1 + c_2 \ln(S - C) + b_2 \text{sample} \times \ln(S - C). \quad (9)$$

The question is whether the interaction term, sample $\times \ln(S - C)$, can be viewed as the same for all samples and this can be answered by using ANCOVA or by multiple regression. We used ANCOVA. If the interaction term of sample $\times \ln(S - C)$ cannot be considered to be the same for all samples, each sample needs a separate regression.

Comparisons across sites were done by selecting a fixed size for diameter or total wet weight. A convenient way of doing this and obtaining 95% confidence limits for gonad weight is by nonlinear regression. In SYSTAT, as part of the estimation of A and β (Eq. 5), an additional line of code can be added as a function of parameters. We used a diameter of 5 cm and total wet weight of 60 g, which are values well within the range of purple sea urchin sizes along the Pacific coast of North America (Ebert 2010a). Using diameter as the covariate

$$\text{FUNPAR G}_5 = A + \beta * \text{LOG}(5 - C) \quad (10)$$

at a diameter of 5 cm or

$$\text{FUNPAR G}_60 = A + \beta * \text{LOG}(60 - C) \quad (11)$$

with total wet weight used as the covariate and 60 g as the fixed size. G_5 and G_60 are the \ln gonad weights for these

sizes. For comparison, analysis was also done using Eq. 4 without a logarithmic transformation because a logarithmic transformation weights small sizes and decreases the contribution to variation of large individuals.

An alternative to using FUNPAR in SYSTAT is to use the linear regressions of \ln gonad weight as a function of $\ln(5 - C)$ for diameter or $\ln(60 - C)$ for total weight and then use the estimated slopes and intercepts to calculate \ln gonad weight as well as the 95% confidence limits for the estimate at 5 cm or 60 g or other value of X (cf. Zar 1974 or any statistics test).

A final problem was how to present information so that it could be compared with other studies. The best approach always is to use the raw data, but much can be done with summary statistics. Our solution is to focus on the minimum number of summary statistics required so results presented here can be used by other investigators to test for differences in gonad development.

Results

Dissection data used to estimate C (Eq. 4) with size, S , as diameter measured in centimeters (Fig. 1) included individuals sufficiently small that a reasonable estimate of C could be obtained. A difficulty was with the dissection of gonads in very small individuals and the resolution of different balances that were available at laboratories. Estimates of C (Table 2) were 0.7 cm for diameter and 0.5 g for wet weight and used with Eqs. 4 and 5 in the following analyses.

An example of data for Gregory Point (Fig. 2) shows seasonal changes from January–December 2007 with

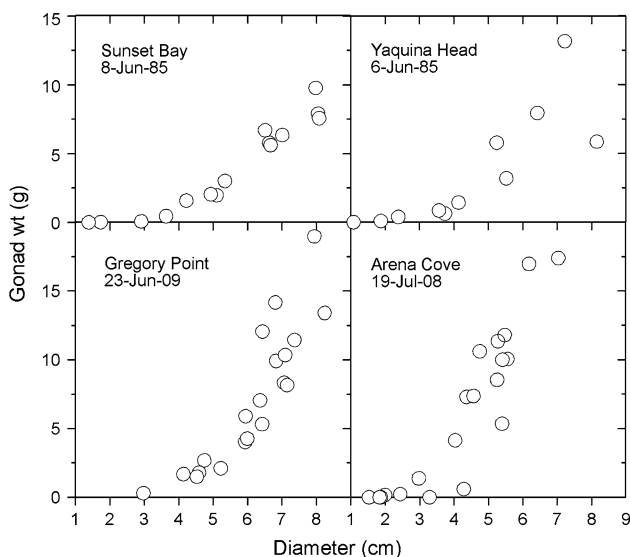


Fig. 1 Four of the samples used for *Strongylocentrotus purpuratus* to estimate C , the diameter when gonads begin to develop (Table 2)

Table 2 Estimates of C , the size when gonads begin to develop

Site	Date	N	R^2	C (cm)	SE
<i>A. Diameter (cm)</i>					
Sunset Bay, OR	8-Jun-85	14	0.97	0.858	0.397
Yaquina Head, OR	6-Jun-85	10	0.92	0.690	1.112
Yaquina Head, OR	17-May-87	10	0.97	0.411	1.053
Gregory Point, OR	16-May-07	20	0.95	0.519	1.088
Gregory Point, OR	23-Jun-09	20	0.93	0.747	1.337
Arena Cove, CA	16-Jan-07	20	0.89	0.858	1.017
Arena Cove, CA	5-May-07	20	0.84	0.627	1.815
Arena Cove, CA	15-Jul-07	20	0.92	0.894	0.435
Arena Cove, CA	18-Jul-08	20	0.76	0.879	0.937
Arena Cove, CA	30-Sep-08	20	0.83	0.869	0.652
Arena Cove, CA	17-Oct-08	20	0.8	0.597	0.408
Arena Cove, CA	25-Nov-08	20	0.89	0.544	0.414
				Mean =	0.708 cm
Site	Date	N	R^2	C (g)	SE
<i>B. Total wet weight (g)</i>					
Boiler Bay, OR	12-Aug and 15-Oct-07	40	0.92	0.499	0.746
Yaquina Head, OR	17-May-87	10	0.97	0.875	2.535
Yankee Point, CA ^a	Dec-65	16	0.98	0.057	0.368
Arena Cove, CA	25-Nov-08	20	0.89	0.455	0.432
				Mean =	0.471 g

^a Based on data from Giese (1967)

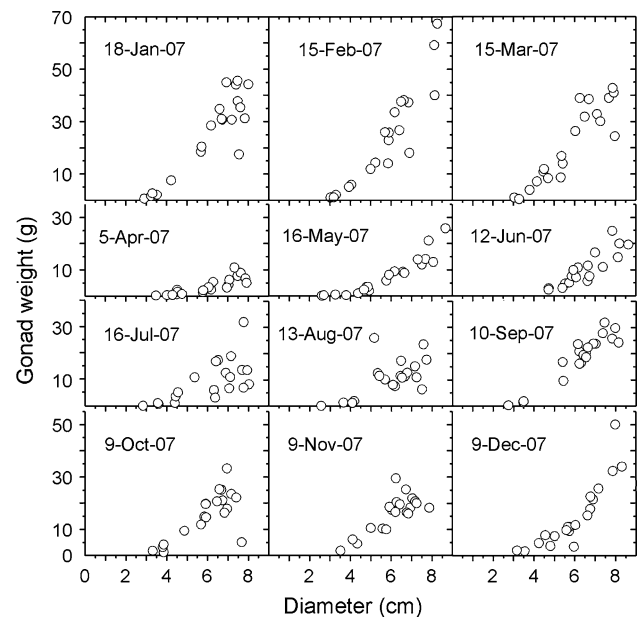


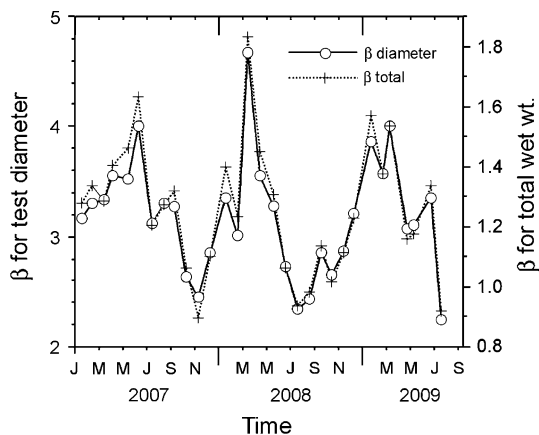
Fig. 2 Monthly dissections of *Strongylocentrotus purpuratus* at Gregory Point, Oregon, showing increasing spread of data points with increasing diameter and the presence of outliers particularly in August and October

Table 3 ANCOVA for monthly gonad dissections of *Strongylocentrotus purpuratus* at Gregory Point, Oregon from January 2007–August 2009; 20 urchins were dissected for each month

Source	SS	df	MS	F-ratio	P
<i>A. Diameter measurements in centimeters</i>					
ln (Dia – 0.7)	449.499918	1	449.4999	3281.373	<0.000001
Sample	20.384785	30	0.67949	4.9603	<0.000001
ln (Dia – 0.7) × Sample	11.826195	30	0.39421	2.8777	0.000001
Error	76.437806	558	0.13699		
Source	SS	df	MS	F-ratio	P
<i>B. Total wet weight in grams</i>					
ln (Total – 0.5)	454.91795	1	454.91795	3321.3185	<0.000001
Sample	21.164651	30	0.705488	5.150712	<0.000001
ln (Total – 0.5) × Sample	13.02493	30	0.434164	3.169798	<0.000001
Error	76.428748	558	0.136969		

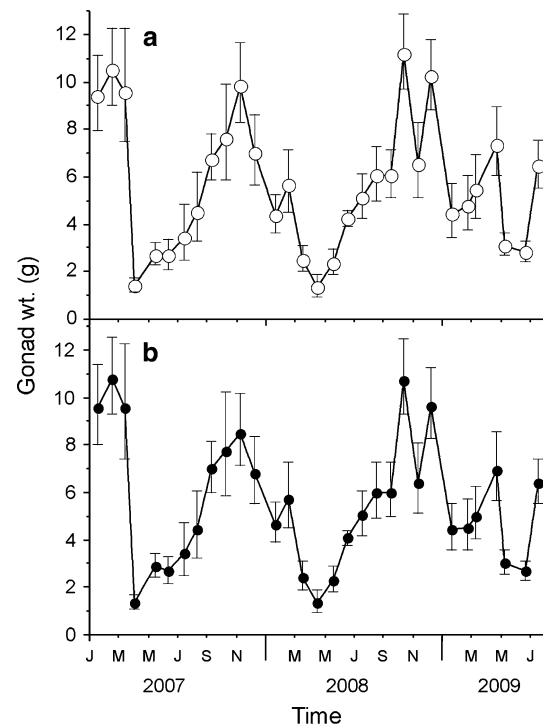
highest values in February and lowest in April. Variation in gonad weights increased with diameter and included outliers such as a high gonad weight in August and a very low gonad weight in October. Outliers were included in all analyses. An ANCOVA for all 31 monthly samples (Table 3) for both independent variables ln(Dia – 0.7) and ln(Total – 0.5) showed that in both cases the interaction of Sample × Independent Variable (diameter or total wet weight) indicated that a common slope was not appropriate for further analysis ($P < 0.00001$). Consequently, individual regressions were used for each sample.

A first approach was to examine slopes for patterns. If just one or two slopes were very different then it still may have been reasonable to continue with a common slope. A pattern of slopes, however, was clear both for ln(Dia – 0.7) and ln(Total – 0.5) (Fig. 3) with high and low points that showed an annual cycle. The largest slope, β , occurred in June 2007, which was 2 months after the smallest gonads in April 2007 (Fig. 2). The smallest slope (Fig. 3) was in

**Fig. 3** Monthly changes of the allometric exponent β with analyses of diameter or total wet weight as covariates and with corrections, C, for when gonads begin to develop

November 2007, 2–3 months before the largest gonads were detected. It is clear (Fig. 3) that use of diameter or total wet weight provide nearly identical patterns and neither provide exponents that consistently correspond with the expected values of 3.0 for diameter or 1.0 for total wet weight.

An approach to comparing sites or samples is to select some size, either diameter (5.0 cm) or total wet weight (60 g), and estimate the gonad wet weight for such a fixed size. This summary, which is virtually identical for diameter or total wet weight (Fig. 4) shows similar patterns of

**Fig. 4** Change in gonad weight for *Strongylocentrotus purpuratus* over 31 months at Gregory Point, Oregon; **a** diameter fixed at 5 cm; **b** total wet weight fixed at 60 g; error bars are 95% confidence limits

peaks and valleys across years but also yearly differences using either diameter or total wet weight with respect to times when gonad weight drops as well as the rate of this change. An interesting feature of the graph is the differences in the 95% confidence intervals, which are largest when gonads are large and smallest when gonads are small. This pattern is not surprising in examining the 2007 data (Fig. 2) where the spread of data points for large individuals is greatest when gonads are large.

The choice of using logarithmic transformation has consequences for how lines pass through the data points because small values are given greater weight in analysis. Differences between transformed and nontransformed data often are not great but in some cases fitted lines can be very different. Data for the first 4 months of the study (Fig. 5) show results for both transformed and untransformed data and for both diameter and total wet weight measurements of the covariate. In all cases using transformed data, a better fit for small individuals is provided and so initially in all cases the line for untransformed data lies above the line for transformed data. In some instances (5 April 2007), relative position remains for all sizes, diameter or total wet weight. For other samples, the lines for transformed and

untransformed cross including going from positive to negative allometry (15 March 2007 total wet weight). The problems lie with how the variance of gonad weight changed with size and if there were outliers, as is the case for the 15 March sample. Outliers always were included in analysis. For purposes of gonad analyses, reducing such variability with a logarithmic transformation is appropriate and is our choice for analysis. Transformation also makes comparison easier because linear models can be used.

Five summary statistics of regressions using Eq. 5 are sufficient to permit other workers to compare their results with ours. These statistics are the means of X and Y , the standard error of the estimate, the coefficient of determination (R^2), and the slope of the regression, which is β . The following steps show how to reconstitute the important products and test for differences of slopes and intercepts. Notation follows Zar (1974). The standard error of the estimate

$$SE_{est} = \sqrt{MS_{resid}} \tag{12}$$

and so

$$SS_{resid} = (SE_{est})^2 \times (n - 2) \tag{13}$$

$$The\ total\ sums\ of\ squares\ SS_{tot} = SS_{resid}/(1 - R^2) \tag{14}$$

and so the sums of squares for the regression

$$SS_{reg} = SS_{tot} - SS_{resid}. \tag{15}$$

To compare samples, one also needs the means for X ($\ln(Dia - 0.7)$ or $\ln(Total - 0.5)$) and Y ($\ln Gonad$). Slopes, β , are compared with a t -test

$$t = \frac{\beta_1 - \beta_2}{s_{\beta_1 - \beta_2}} \tag{16}$$

where

$$s_{\beta_1 - \beta_2} = \sqrt{\frac{(s_{y \cdot x}^2)_p}{(\sum x^2)_1} + \frac{(s_{y \cdot x}^2)_p}{(\sum x^2)_2}}. \tag{17}$$

The pooled residual mean square is

$$(s_{y \cdot x}^2)_p = \frac{(SS_{resid})_1 + (SS_{resid})_2}{(DF_{resid})_1 + (DF_{resid})_2}. \tag{18}$$

It also is necessary to know the sums of squares of x and cross products xy , for each sample,

$$\sum x^2 = \frac{(\sum xy)^2}{SS_{reg}} \tag{19}$$

and

$$\sum xy = \frac{SS_{reg}}{b}. \tag{20}$$

The calculations for t can now be completed and the test has $n_1 + n_2 - 4$ degrees of freedom.

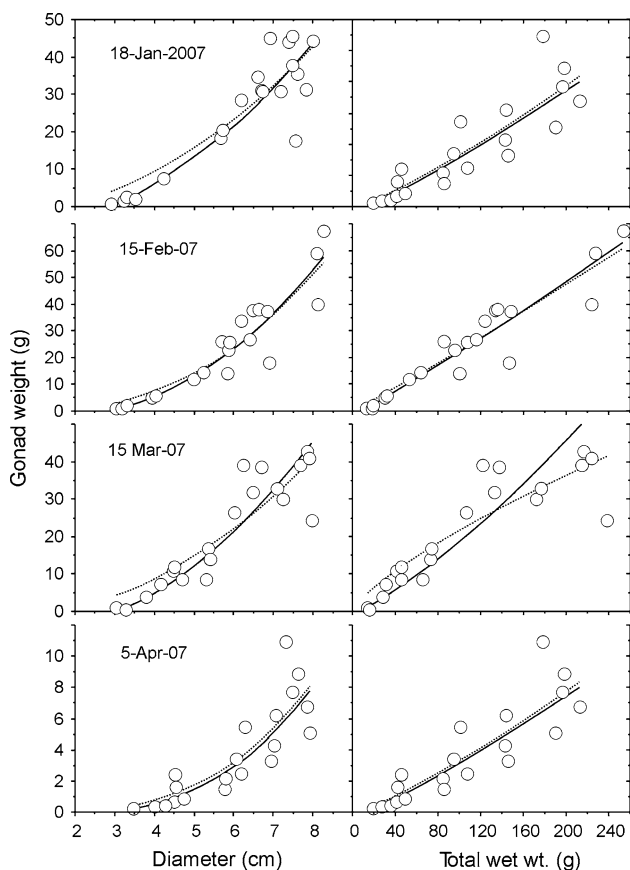


Fig. 5 Examples of fitted lines for *Strongylocentrotus purpuratus* gonad data using transformed (solid line Eq. 5) and untransformed (dashed line Eq. 4) data and both diameter and total wet weight

If slopes are judged to be different, this is the end of the analysis and separate regressions would be used for the two samples that include different slopes and intercepts. If, on the other hand, slopes are judged to be equal, analysis proceeds to a test for equality of intercepts. The first step is calculation of a common or weighted slope

$$\beta_c = \frac{(\sum xy)_1 + (\sum xy)_2}{(\sum x^2)_1 + (\sum x^2)_2} \tag{21}$$

The calculation of a *t*-value also requires a residual sum of squares for the common regression

$$\begin{aligned} (s_{y \cdot x}^2)_c &= \frac{(SS_{tot})_1 + (SS_{tot})_2 - ((\sum xy)_1 + (\sum xy)_2)^2 / ((\sum x^2)_1 + (\sum x^2)_2)}{n_1 + n_2 - 3} \end{aligned} \tag{22}$$

and so

$$t = \frac{(\bar{Y}_1 - \bar{Y}_2) - b_c(\bar{X}_1 - \bar{X}_2)}{\sqrt{(s_{y \cdot x}^2)_c \left[\frac{1}{n_1} + \frac{1}{n_2} + \frac{(\bar{X}_1 - \bar{X}_2)^2}{(\sum x^2)_1 + (\sum x^2)_2} \right]}} \tag{23}$$

with $n_1 + n_2 - 3$ degrees of freedom. An EXCEL file is available as an electronic supplement (Online Resource 1) to do these calculations with the summary statistics.

A comparison of results using the summary statistics compared with raw data for the 12 June 2007 from Gregory Point and the 8 June 1985 sample from Sunset Bay (Table 4) shows that the results are identical. Table 4B first shows the test for the equality of slopes by a *t*-test followed by the test for equality of intercepts. Table 4C is an ANCOVA (GLM, SYSTAT 2004) using raw data first testing for the significance of the interaction term of Sample \times ln(Dia - 0.7) followed by a repeat of the analysis without interaction. The *P*-values are the same for the two analyses.

For comparison with analyses using Equations 4 and 5, a mean gonad index (Eq. 1) was calculated for each of the 31 monthly samples from Gregory Point using gonad and total wet weight (Fig. 6). The annual cycle is clear and shows differences across years and so discerning a pattern was not sensitive to violating the assumptions of the allometric exponent $\beta = 1.0$ and the correction for gonad development starting at $C = 0$. The gonad index, GI, was always above the other two measures, which is reasonable because the monthly total weights ranged from 110 to 131 g and so were always greater than 60 g. Differences, however, were greatest when gonads were large and smallest when gonads were small such as in April 2007 and 2008. An important point is that comparing the GI results with other studies or across years cannot be done because of the allometric problems presented here. The differences

Table 4 Comparison of samples from Gregory Point and Sunset Bay using summary statistics (A and B) and ANCOVA using raw data (C)

Sample	<i>N</i>	Mean <i>X</i>	Mean <i>Y</i>	<i>R</i> ²	SE _{est}	Slope <i>b</i>
<i>A. Summary statistics</i>						
Greg. Pt 12-Jun-07	20	1.679126	1.868634	0.928457	0.362400	4.000330
SS Bay 8-Jun-85	14	1.482973	0.612659	0.965010	0.388767	3.547010
Test	<i>t</i>	<i>df</i>	<i>P</i>			
<i>B. Analysis</i>						
Equality of slopes	1.3816	30	0.1773	Common slope = 3.694488		
Equality of elevations	3.9230	31	0.0005			
Source	SS	<i>df</i>	MS	<i>F</i> -ratio	<i>P</i>	
<i>C. ANCOVA with raw data</i>						
First a test of the interaction term Sample \times ln(Dia - 0.7)						
ln(Dia - 0.7)	73.6772	1	73.6772	529.0769	<0.0001	
Sample	0.0188	1	0.0188	0.1352	0.7157	
Sample \times ln(Dia - 0.7)	0.2658	1	0.2658	1.9087	0.1773	
Error	4.1777	30	0.1393			
The <i>P</i> -value indicates that slopes should be considered equal so analysis is repeated without the interaction term						
Sample \times ln(Dia - 0.7) to test for equality of elevations						
ln(Dia - 0.7)	80.4339	1	80.4339	561.1475	<0.0001	
Sample	2.2060	1	2.2060	15.3899	0.0005	
Error	4.4435	31	0.1433			

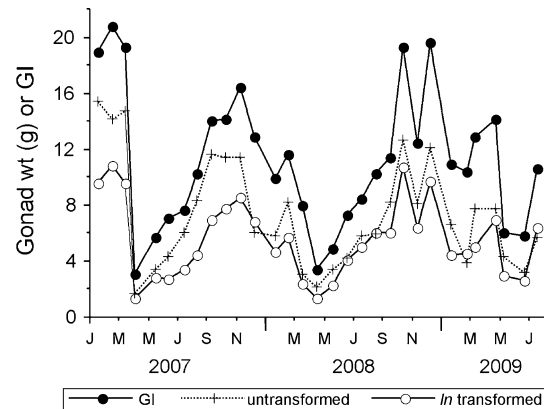


Fig. 6 Comparison of the gonad index, GI, with gonad weights of 60 g individuals using ln-transformed (Eq. 5) and untransformed (Eq. 4) data for analysis

shown by comparing the GI measures with a fixed size at 60 g illustrate what would happen if the Gregory Point data would be compared with another site where urchins were smaller but had gonad development that was the same as at

Table 5 Summary statistics for Gregory Point necessary to make comparisons with other data sets

ln(Dia – 0.7)						ln(Total – 0.5)			
Date	Mean X	Mean Y	R ²	SE _{est}	β	Mean X	R ²	SE _{est}	β
18-Jan-07	1.623894	2.780484	0.932627	0.345794	3.163906	4.508816	0.939573	0.327483	1.278657
15-Feb-07	1.561315	2.716974	0.941687	0.306905	3.309339	4.358372	0.938885	0.314193	1.334371
15-Mar-07	1.567001	2.618474	0.845128	0.509876	3.334608	4.364538	0.851047	0.500038	1.286367
5-Apr-07	1.623559	0.863748	0.859769	0.430335	3.552174	4.441182	0.873723	0.408364	1.403086
16-May-07	1.533483	1.327920	0.951083	0.348670	3.524766	4.312382	0.951604	0.346808	1.460516
12-Jun-07	1.679126	1.868634	0.928457	0.362400	4.000330	4.629693	0.913495	0.398497	1.631903
16-Jul-07	1.627732	1.757484	0.777751	0.622746	3.115279	4.514899	0.760028	0.647099	1.202826
13-Aug-07	1.606500	1.977332	0.782872	0.603856	3.311147	4.454059	0.780529	0.607105	1.273040
10-Sep-07	1.658567	2.599026	0.946835	0.277084	3.273598	4.608798	0.953990	0.257766	1.316229
9-Oct-07	1.614310	2.461821	0.700853	0.520154	2.639975	4.490030	0.740622	0.484346	1.060202
9-Nov-07	1.660099	2.636616	0.816147	0.287834	2.458353	4.480116	0.789045	0.308320	0.897178
9-Dec-07	1.610656	2.351104	0.841512	0.385011	2.851877	4.456740	0.822600	0.407334	1.100167
20-Jan-08	1.683387	2.293473	0.921631	0.296994	3.353287	4.669933	0.925336	0.289889	1.400881
18-Feb-08	1.700668	2.479511	0.844502	0.384580	3.017307	4.689144	0.857258	0.368468	1.231660
17-Mar-08	1.680871	1.915570	0.904342	0.414046	4.677095	4.633069	0.926954	0.361814	1.834413
16-Apr-08	1.730519	1.249900	0.726963	0.454827	3.551326	4.742478	0.761602	0.424998	1.451737
20-May-08	1.626549	1.375538	0.820471	0.421331	3.279435	4.485397	0.830888	0.408925	1.306484
18-Jun-08	1.661508	1.961731	0.976128	0.131924	2.730618	4.576119	0.974952	0.135133	1.062926
17-Jul-08	1.693622	2.164595	0.781410	0.268673	2.345807	4.658926	0.796040	0.259526	0.936657
17-Aug-08	1.710172	2.403011	0.820292	0.275615	2.436287	4.701094	0.828703	0.269088	0.983299
16-Sep-08	1.708334	2.509199	0.860947	0.276888	2.857314	4.708833	0.892721	0.243204	1.136685
14-Oct-08	1.713886	3.050741	0.892622	0.191308	2.650083	4.711072	0.892175	0.191706	1.016754
11-Nov-08	1.709409	2.577223	0.842714	0.358678	2.868121	4.718490	0.824848	0.378502	1.114566
9-Dec-08	1.650806	2.886705	0.913871	0.259206	3.219133	4.544310	0.919955	0.249882	1.227161
22-Jan-09	1.639816	2.189413	0.906680	0.411383	3.858299	4.528638	0.878260	0.469867	1.567373
23-Feb-09	1.660852	2.232215	0.842950	0.388576	3.575993	4.570397	0.843234	0.388225	1.373697
10-Mar-09	1.647363	2.364587	0.908927	0.393975	3.991557	4.524235	0.878093	0.455815	1.535567
24-Apr-09	1.724087	2.755876	0.865167	0.288380	3.079650	4.742094	0.870894	0.282189	1.160568
10-May-09	1.722779	1.920245	0.905231	0.237335	3.103343	4.753133	0.921787	0.215609	1.176879
23-Jun-09	1.647015	1.613116	0.934491	0.268021	3.353655	4.514755	0.932933	0.271190	1.334615
20-Jul-09	1.709943	2.421587	0.883836	0.224695	2.247265	4.688953	0.867086	0.240349	0.922270

Mean of Y is ln gonad wet weight and mean of X is for ln(Dia – 0.7 cm) and ln(Total wet weight – 0.5). Natural logarithms are used. The large number of significant figures are important when they are used for comparison with other data sets

Gregory Point. Sea urchins at Gregory Point would be judged to be more productive whereas they really would be just larger individuals.

Discussion

We have shown that simple assumptions about the allometry of gonad development in a sea urchin make the gonad index, GI (Eq. 1) suspect for any application other than description of the reproductive cycle, which has been recognized by others (Marsh and Watts 2007). Because gonads begin to develop at a test diameter greater than

zero, an additional parameter, C , is required in analysis. Estimation of this parameter from dissection data has large standard errors (Table 2) and so the values presented here for *S. purpuratus* must be viewed as provisional. The estimate of $C = 0.7$ cm for test diameter when gonads begin to develop, however, is similar to the estimates of 0.7–0.8 cm for *S. droebachiensis* (O. Ellers, personal communication). To our knowledge, there are no histological studies that would provide clearer evidence for when gonads in *S. purpuratus* begin to develop.

Data from Gregory Point, Oregon, showed that the allometric exponent β was not usually equal to 1.0 when total wet weight was used as the independent variable or

equal to 3.0 when diameter was used (Fig. 3). Furthermore, the exponent changed during the year so a single value could not be used. Individual regressions were required to describe each sample (Table 5). If a research goal includes comparing size-specific gonad development or differences between sites or times (e.g., Lessios 1981; Boivin et al. 1986; Lawrence and Byrne 1994; Lozano et al. 1995; Lester et al. 2007), use of the simple GI runs the risk of producing errors that may be serious (c.f. Beupre and Dunham 1995).

A solution we have presented here is to compare samples by selecting a specific size, either diameter or total weight, and estimating gonad size although other measures of total and gonad size could be used such as volume or dry weights. We selected 5 cm test diameter and 60 g total wet weight because these are values found in samples over a very wide geographic range for *Strongylocentrotus purpuratus*; individuals with diameters ≥ 8 cm are not present at most sites along the coast (Ebert 2010a). Comparing samples in other studies can be done using the five summary statistics we have provided; statistical comparisons of gonad index, GI, values across studies cannot be done when different sizes of individuals are being compared and, given the problems presented here, should not be done within a study.

Using a wide size range of individuals in a study has further benefits. Comparisons of life histories of sea urchins are difficult. In the past, when the Brody–Bertalanffy growth model still seemed appropriate (Ebert 1975), a comparison of the parameter K and the instantaneous mortality rate, M , was reasonable. Indeed, if growth data are forced to fit the Brody–Bertalanffy equation (Ebert 2007) comparisons still can be done in this manner. If a different growth model is used, however, such as the Richards or Tanaka, there no longer is a single parameter that can represent growth. An alternative is to use production and biomass estimates (Ebert 2010b), an approach that requires estimating age-specific spawn biomass. Gonad data gathered for all available sizes and at times of maximum and minimum gonad size can be used to calculate size-specific spawn mass and, with growth estimates, age-specific spawn mass (e.g., Ebert 2008).

Our recommendations for studies where body components are to be compared are as follows:

1. Gather a wide range of sizes including small individuals rather than concentrating on a restricted size range. Work of dissection is the same and information content is much greater.
2. Use the allometry model with an adjustment for size when gonads begin to develop (Eq. 4).
3. Publish the summary statistics.
4. Save the data in a form that can be shared because there is no substitute for the raw data.

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